

Evaluation of a new supply strategy based on stochastic programming for a fashion discounter

Miriam Kießling, Tobias Kreisel, Sascha Kurz and Jörg Rambau

Abstract Fashion discounters face the problem of ordering the right amount of pieces in each size of a product. The product is ordered in pre-packs containing a certain size-mix of a product. For this so-called lot-type design problem, a stochastic mixed integer linear program was developed, in which price cuts serve as recourse action for oversupply. Our goal is to answer the question, whether the resulting supply strategy leads to a supply that is significantly more consistent with the demand for sizes compared to the original manual planning. Since the total profit is influenced by too many factors unrelated to sizes (like the popularity of the product, the weather or a changing economic situation), we suggest a comparison method which excludes many outer effects by construction. We apply the method to a real-world field study: The improvements in the size distributions of the supply are significant.

1 Introduction

The lot-type design problem LDP seeks for an optimal set of lot-types and a supply in terms of lots of such lot-types such that the resulting supply of sizes matches the branch- and size-dependent demand as closely as possible (see [1] for details). A *lot-type* is defined as a vector with a component for each size. This component specifies the number of pieces of that size in the lot. For example, for the sizes S , M , L , and XL , a lot of lot-type $(1, 2, 3, 1)$ contains one item of size S , two of size M , three of size L , and one of size XL .

In reality, excess supply is compensated by cutting prices. Therefore, we extended the model in [1] by a model for this recourse action. This resulted in the following stochastic mixed integer linear program, in the following denoted by SLDP:

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$$\max - \sum_{b \in B} \sum_{l \in L} \sum_{m \in M} c_{b,l,m} x_{b,l,m} - \sum_{i=1}^n \delta_i \cdot z_i + \mathbb{E}_{\xi}(Q(x, \xi)) \quad (1)$$

$$\text{s. t.} \quad \sum_{l \in L} \sum_{m \in M} x_{b,l,m} = 1 \quad \forall b \in B, \quad (2)$$

$$\underline{I} \leq \sum_{b \in B} \sum_{l \in L} \sum_{m \in M} m \cdot |l| \cdot x_{b,l,m} \leq \bar{I}, \quad (3)$$

$$\sum_{l \in L} y_l \leq \sum_{i=1}^n z_i, \quad (4)$$

$$\sum_{m \in M} x_{b,l,m} \leq y_l \quad \forall b \in B, l \in L, \quad (5)$$

$$z_i \leq z_{i-1} \quad \forall i \in \{2, \dots, n\}, \quad (6)$$

$$x_{b,l,m} \in \{0, 1\} \quad \forall b \in B, l \in L, m \in M, \quad (7)$$

$$y_l \in \{0, 1\} \quad \forall l \in L, \quad (8)$$

$$z_i \in \{0, 1\} \quad \forall i \in \{1, \dots, n\}, \quad (9)$$

The meanings of the symbols are as follows: By L we denote the set of possible lot-types which can be delivered to a branch from the set B with multiplicity from the set M . If Branch b gets Lot-type l with Multiplicity m then the corresponding binary variable $x_{b,l,m}$ takes value 1, and 0 otherwise. At most n lot-types may be used. If we deliver Lot-type l to Branch b with Multiplicity m costs of $c_{b,l,m}$ arise. Every branch is supplied by exactly one lot-type with one multiplicity, see constraint (2). The overall supply has to be between the lower bound \underline{I} and the upper bound \bar{I} (3). By $|l|$ we denote the overall number of items in lot-type l . Using i lot-types implies costs of δ_i . Constraint (6) requires that also the costs for using $i-1, i-2, \dots, 1$ are added if i different lot-types are used. Due to Constraint (5), y_l indicates whether some branch $b \in B$ is supplied by Lot-type l . Finally, (4) links y_l and z_i .

To compensate for excess supply – depending on a vector of random variables ξ describing the demand – price cuts are possible. The corresponding optimization problem seeking for optimal price-cut strategies is denoted by $Q(x, \xi)$ where x is the vector of all variables $x_{b,l,m}$.¹ The objective function is the revenue depending on supply x and the random vector ξ . Our goal is to maximize the total profit in expectation.

We performed a field study at our project partner using the SLDP. How can we find out, whether the new method outperforms the traditional manual planning? From an economic point of view, we should look at the actual revenues. However, the data showed that actual revenues during the field study are distorted too much by factors beyond our control. For this purpose we developed two different comparison methods which exclude such extern distortions and just reveal how well the size-dependent demand is met. We examine significance by using the Wilcoxon rang sum test from statistics.

¹ An approximation of $Q(x, \xi)$ can be computed by solving a mixed integer linear programm or a dynamic program.

In the following we introduce our approaches and show results from the field study.

2 Comparison methods

We developed two different indicators for demand consistency of the supply with sizes.

The first approach – called Normalized Sales Rate Deviation (NSRD) – compares the sales rates per size against each other in such a way that the popularity of the product itself has no dominating influence on the result anymore.

The simple and course idea is to observe the sales at the 50 %-day and look for each size at the fraction of supply that has been sold so far. The 50 %-day is the first day in the sales period where 50 % of the total supply of the product has been sold. We get an estimation of a normalized sales rate of each size relative to its supply, independent of the popularity of the product. We call this number the *normalized sales rate estimate* of a size, denoted by $\widehat{\text{NSR}}(s)$.

For example, if there is a product with supply $(10, 20, 20, 10)$, and the size-dependent sales numbers up to the 50 %-day are $(2, 10, 15, 3)$, then the estimated sales rates are $(0.2, 0.5, 0.75, 0.3)$. This indicates that the supply of M was spot-on (i.e., relative sales in this size were the same as relative sales in total), whereas the supplies of S and XL were too small and the supply of L was too large in the considered branches.

Based on these observations, we are interested in how much the sales in a size deviate from the overall 50 %. To this end, we estimate the standard deviation of the observed normalized selling rates (relative to the supply). Because of multiple sales per day we can have an average NSR different from 0.5. Therefore, the standard deviation must be taken with respect to the sample average, which is²

$$\text{SD}^+ := \sqrt{\frac{1}{N-1} \sum_{i=1}^N (r_i - \hat{r})^2}, \quad (10)$$

where N is the sample size and \hat{r} the sample average. In our case, we compute for N sizes

$$\text{NSRD} := \sqrt{\frac{1}{N-1} \sum_{i=1}^N (\text{NSR}(s_i) - \widehat{\text{NSR}})^2}, \quad (11)$$

the *normalized sales rate deviation* (for a single product).

The smaller the normalized sales rate deviation is, the more consistent is the supply with the demand for sizes.

The second approach uses ideas from [5]. Let P be the set of products in a given commodity group, S_p be the set of sizes for a product $p \in P$, b be a given branch,

² We use the notation from [3].

s be a given size, and $\theta_{b,s'}(p)$ be the first day at which Product p is sold out in Branch b and Size s' (here $\theta_{b,s'}(p) = \infty$ is possible). With this we can define the Top-Dog-Count $W(b, s)$ as

$$\left| \left\{ p \in P \mid 0 = |\{s' \in S_p \mid \theta_{b,s'}(p) < \theta_{b,s}(p)\}| \right\} \right| \quad (12)$$

and the Flop-Dog-Count $L(b, s)$ as

$$\left| \left\{ p \in P \mid 0 = |\{s' \in S_p \mid \theta_{b,s'}(p) > \theta_{b,s}(p)\}| \right\} \right|. \quad (13)$$

We now consider the value $|W(b, s) - L(b, s)|$, in the following denoted as *top-dog-deviation of Size s* , denoted by $\text{TDD}(b, s)$. We assume that a value closer to zero corresponds to a better supply strategy for s , because $\text{TDD}(b, s)$ estimates the difference of the probability that s is sold out first and the probability that s is sold out last. In order to exclude numerical artefacts with too small integers, we restrict ourselves to samples that consist of a (random) subset of articles such that $W(b, s) + L(b, s)$ equals a given number. For each such sample we compute $\text{TDD}(b, s)$.

Since for each branch and each size we get an individual measurement, the TDD is a finer measurement than the NSRD.

Now, any improvement measured by the above indicators in the field study could be a mere coincidence. To obtain statements about the significance of the observations NSRD and $\text{TDD}(b, s)$, we apply the non-parametric *Wilcoxon rank-sum test* (see [2]). This method tests whether two sets of realizations stem from the same distribution.

Let us first sketch the Wilcoxon-rank-sum-test. Given two sets of realizations A and B , the null hypothesis is that A and B stem from the same distribution. The alternative says that distribution of A is shifted to the left.

In our context the alternative means: there is less deviation among normalized sales rates in the distribution of A than there is in the distribution of B . Or, respectively, A has smaller $\text{TDD}(b, s)$ s than B .

The test is then conducted by sorting the values of both samples in ascending order, thereby assigning a rank to each value. In the next step the sum of the ranks for Family A is computed, which yields an observed rank sum.

In order to estimate the probability that the distribution of A is shifted to the left compared to the distribution of B , we have to consider the probability for getting rank sums smaller than or equal to the observed rank sum for A . If this probability lies below a predetermined significance level (5% or 10%) we say that the observed difference can not be explained by chance, and we can assume that the A 's distribution is shifted to the left compared to the distribution of B .

3 Comparison results

We employed the SLDP on results of a field study with ladieswear shirts with sales periods from February to June 2011. For comparison we took historical data from the same commodity group in a time period from March until December 2006 – at this time still all items were supplied by manual planning. We consider 26 branches and four different sizes, namely S, M, L and XL.

For the historical sample set only one lot-type was delivered, namely (1, 2, 2, 1). In the field study our system provided the lot-types (1, 1, 1, 1), (1, 1, 2, 2) and (2, 2, 3, 4). Now we would like to find out whether SLDP led to supplies that were significantly more consistent with the demand than the supplies suggested by manual planning.

At first we have a look at the normalized sales rate deviation. Aggregating over the tested branches and rounding to two decimal figures leads us to the observations for NSRD with the corresponding ranks in braces in Table 1.

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|------------------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|------------|-----------|------------|-----------|-----------|----------|------------|----------|------------|------------|------------|------------|------------|------------|
| SLDP | 8.58 (4) | 16.92 (18) | 13.49 (13) | 17.32 (21) | 17.29 (20) | 14.37 (14) | 12.83 (12) | 9.85 (5) | 6.61 (3) | 16.40 (17) | | | | | | | | | | | | | | | | | | |
| manual planning | 18.15 (24) | 21.89 (28) | 22.71 (31) | 12.44 (9) | 19.93 (26) | 22.98 (33) | 27.46 (35) | 21.51 (27) | 22.79 (32) | 15.78 (15) | 12.49 (10) | 16.09 (16) | 22.62 (30) | 19.09 (25) | 17.06 (19) | 10.52 (6) | 17.53 (22) | 11.66 (7) | 12.35 (8) | 5.09 (2) | 17.65 (23) | 3.37 (1) | 22.08 (29) | 12.75 (11) | 23.92 (34) | 27.52 (36) | 42.32 (38) | 36.80 (37) |

Table 1 standard deviation of selling rates (NSRD) for tested articles of ladieswear with ranks in braces

We get a mean of 13.36 for the SLDP and 19.16 for the manual planning. So in average we get smaller deviations among the normalized sales rates per size for SLDP. Thus, we conclude that the systematic error in size distribution is smaller. To test significance, we apply the Wilcoxon rank-sum test with a predefined significance level of 5%. The null hypothesis is that the distributions of NSRD for SLDP and the manual planning are identical, the alternative that for SLDP the distribution is shifted to the left. The test yields a rank sum of 127 for the SLDP. The probability of a rank sum lower than or equal to 127 is approximately 1.17%, which is below the significance level. In other words: The probability that the observed improvements are a coincidence is well below the significance level 5%, and the observations are significant.

As a next step we check the measurements $TDD(b, s)$. Analogous to the selling rates, we compute mean values over all products, branches, and sizes, denoted by \widehat{TDD} . For the SLDP, \widehat{TDD} is 2, whereas it is 3 for the manual planning. Again, we perform the Wilcoxon rank-sum test with a significance level of 5%. We get the following result: With 208 observations (104 for each sample set) and a rank sum of 8766 for the SLDP method, the probability for a shift to the left of the distribution of \widehat{TDD} for SLDP method is less than $10^{-6}\%$. This also is well below the significance level: The observed improvements are significant.

4 Conclusion

We presented methods to compare the impact of different supply strategies for a fashion discounter in a real-world field study. The normalized sales rate deviation among sizes measures how evenly a product sells in the various sizes. The top dog deviation measures to what extent a size is sold out first more often than last, or vice versa. Both measures showed that our new supply strategy based on the stochastic lot design problem can significantly improve the demand consistency of the supply with respect to sizes, where significance is checked by the Wilcoxon rank sum test. Since the Wilcoxon test is robust and uses no assumptions on the distributions, we are confident that this result is practically relevant.

References

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